## Robust Two Sample Analysis

The JAGS analog to the two-sample t-test follows directly from the format described in $\mathbf{0 1 0} \boldsymbol{\&} 020$ MCMC Worksheets. Scaffolds are derived from J.K. Kruschke (K): Doing Bayesian Data Analysis - A Tutorial with R, JAGS, and Stan, available at https://sites.google.com/site/doingbayesiandataanalysis/. The only difference in One-sample versus Two-sample analysis is specifying a second variable in JAGS, derived from a column in the original data, that identifies group membership in the standard long-form data format. In K's scaffold, information about group column name is passed from the Driver level to the function genMCMC() at the 2nd level which in turn codes a numeric variable for JAGS.
from; Jags-Ymet-Xnom2grp-RrobustHetDRIVER

```
yName="Score"
xName="Group"
# Load the relevant model into R's working memory:
source("Jags-Ymet-Xnom2grp-MrobustHet.R")
#----------------------------------------------------------------------------
# Generate the MCMC chain:
mcmcCoda = genMCMC( datFrm=myDataFrame, yName=yName , xName=xName ,
    numSavedSteps=50000 , s aveName=fileNameRoot )
```



```
from: Jags-Ymet-X2grp-MrobustHet.R
\# THE DATA.
    \(y=\) as.numeric(datFrm[,yName])
    \(\mathrm{x}=\) as.numeric(as.factor(datFrm [,xName]))
    xLevels \(=\) levels(as.factor(datFrm[,xName]))
    Ntotal = length \((y)\)
\#-
    \# THE MODEL.
    modelString \(="\)
    model \{
    for (i in 1:Ntotal) \{
        \(y[i] \sim \operatorname{dt}\left(\operatorname{mu}[x[i]], 1 / \operatorname{sigma}[x[i]]^{\wedge 2}\right.\), nu \()\)
    \}
    for ( j in 1:2) ) \# 2 groups
        mu[j] ~ dnorm ( meanY, 1/(100*sdY)^2 )
        sigma[j] ~ dunif( sdY/1000 , sdY*1000 )
    \}
    \(n u \sim \operatorname{dexp}(1 / 30.0)\)
\}
    " \# close quote for modelString
```

The Two-sample Robust model structure looks identical in structure to the One-sample model. The difference lies in addition of index $\boldsymbol{j}$, with values of 1 or 2 , specifying group membership. This means that for the distribution of $y_{i}$, there are two mean parameters $\mu_{1}$ and $\mu_{2}$ and two precision parameters $\tau_{1}$ and $\tau_{2}$ for each group. Priors for each of the four distribution parameters for $y_{i}$ are carried out independently. There is, however, only one normality parameter $v$, so only a single prior is detemined for this exactly as in the One-sample model.

## Two Sample t-tests:

For comparison with MCMC resuls, the following are results from running standard t-tests with and without assuming equality of variance between the two group:

```
> t.test(Score~Group,data=myDataFrame, var.equal=FALSE) # unequal variances
            Welch Two Sample t-test
data: Score by Group
t = -1.958, df = 111.44, p-value = 0.05273
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
    -15.70602585 0.09366161
sample estimates:
            mean in group Placebo mean in group Smart Drug
                100.0351 107.8413
> t.test(Score~Group,data=myDataFrame, var.equal=TRUE) #equal variances
            Two Sample t-test
data: Score by Group
t = -1.9249, df = 118, p-value = 0.05665
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
    -15.8369851 0.2246208
sample estimates:
            mean in group Placebo mean in group Smart Drug
                        100.0351 107.8413
> var.test(Score~Group,data=myDataFrame)
            F test to compare two variances
data: Score by Group
F = 0.49457, num df = 56, denom df = 62, p-value = 0.00823
alternative hypothesis: true ratio of variances is not equal to 1
95 percent confidence interval:
    0.2962264 0.8313852
sample estimates:
ratio of variances
                        0.4945694
> sd(myDataFrame$Score[myDataFrame$Group=="Placebo"])
[1] 17.8945
> sd(myDataFrame$Score[myDataFrame$Group=="Smart Drug"])
[1] 25.4452
```

The variance ratio F-test (see Biostatistics 170) indicates that the Welsh Test above is more appropriate.

## Two Sample MCMC Results:

| > show(summarylnfo) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean |  | Median | Mode | ESS | HDImass |
| mu [1] | 99.2601668 | 99.27 | 72400099. | . 2606903 | 30169.3 | 0.95 |
| mu [2] | 107.1396293 | 107.14 | 440000107. | . 0112339 | 29957.3 | 0.95 |
| muDiff | 7.8794625 | 7.88 | 8380007. | . 9411225 | 29253.3 | 0.95 |
| sigma[1] | 11.3367363 | 11.21 | 15400011. | . 2390852 | 11999.9 | 0.95 |
| sigma[2] | 17.9232785 | 17.77 | 79500017. | . 4576198 | 11537.8 | 0.95 |
| sigmaDiff | 6.5865422 | 6.50 | 0730006. | . 4381478 | 22301.2 | 0.95 |
| nu | 3.8650320 | 3.50 | 0613002. | . 9775310 | 5588.0 | 0.95 |
| $\operatorname{log10}(\mathrm{nu})$ | 0.5570688 | 0.5 | 448280 0. | . 5417733 | 7816.5 | 0.95 |
| effsz | 0.5319058 | 0.52 | 2916140. | . 5007781 | 26530.7 | 0.95 |
|  | HDIlow | HDIhigh CompVal PentGtCompVal |  |  |  |  |
| mu [1] | 95.79930000 | 102. | 9110000 | NA | NA |  |
| $m u[2]$ | 101.78200000 | 112.3780000 |  | NA |  | NA |
| muDiff | 1.69020000 | 14.3730000 |  | 0 | 99.20802 |  |
| sigma[1] | 8.08601000 | 14.7977000 |  | NA | NA |  |
| sigma[2] | 12.84890000 | 23.3849000 |  | NA | NA |  |
| sigmaDiff | 1.16020000 | 12.0011100 |  | 0 | 99.31401 |  |
| nu | 1.61057000 | 6.8828200 |  | NA | NA |  |
| $\operatorname{log10(nu)~}$ | 0.26902467 | 0.8689435 |  | NA | NA |  |
| effsz | 0.08978173 | 0.9767674 |  | 0 | 99.20802 |  |
|  | ROPElow ROPEhigh |  | PcntLtROPE | PcntInROPE PcntGtROPE |  |  |
| mu [1] | NA | NA | NA |  | NA | NA |
| mu [2] | NA | NA | NA |  | NA | NA |
| muDiff | -0.5 | 0.5 | 0.5079898 | 0.68398 | 86398 | . 80802 |
| sigma[1] | NA | NA | NA |  | NA | NA |
| sigma[2] | NA | NA | NA |  | NA | NA |
| sigmaDiff | -0.5 | 0.5 | 0.4139917 | 0.67798 | 86498 | . 90802 |
| nu | NA | NA | NA |  | NA | NA |
| $\operatorname{log10}(\mathrm{nu})$ | NA | NA | NA |  | NA | NA |
| effsz | -0.1 | 0.1 | 0.2019960 | 2.31795 | 53697 | . 48005 |

## Summary of Findings:

The Variance Ratio test shows strong preference for the two groups having different variances ( $\mathbf{p}$-value $=\mathbf{0 . 8} \%$ ). Similar results are seen in the MCMC plot for Differences in Scales ( $0.7 \%$ for zero difference in the mode and outside the HDI). The MCMC Post Predicion Plot shows a difference quite clearly both in the distribution of the original data (red histogram) and the different shapes of the probable t-distributions for each group. The fit looks good. The Normality plot with mode $=0.542$, quite a bit less than $\log _{10}(30)=1.4771$, indicates preference for the $t$-distribution versus Normal distribution as a better fit for the data, given the evident data outliers. As a result of using the $t$-distribution, lower variance is seen in the fit of MCMC compared with measurement of standard deviation fron the samples. MCMC difference in modes of $\mu$ includes only $0.8 \%$ of the distribution and outside the $95 \%$ HDI - impressive) whereas Welsh p-value for difference in means is $0.0523 \mathbf{( 5 . 2 3 \%}$ - not significant).

## How do the Welsh/Variance Ratio tests and MCMC Results compare?

| Parameter: | Welsh Test: | 95\% CI |
| :---: | :---: | ---: |
| $\mu$ | sample means: | for the difference: |
|  | Placebo $=100.0351$ | $[0.0934-15.706]$ |

$\sigma \quad$ sample sd:
Placebo $=17.8945$
Smart Drug $=25.4452$
Ratio $=0.4946$
for the ratio:
[0.2962-0.8314]

MCMC using JAGS: modes:
Placebo $=99.3$
Smart Drug = 107
modes:
Placebo = 11.2
for the difference:
[1.16-12]



Correlation of $v$ with both shape parameters $\sigma_{1}$ and $\sigma_{2}$ is to be expected from the definition of the $\mathbf{t}$-distribution. Shape parameters $\sigma_{1}$ and $\sigma_{2}$ are show some correlation with each other. Otherwise correlations are low.

