$ORIGIN \equiv 1$

ANOVA Randomized Block Design

ANOVA designs involving "randomized blocks with replicates", also termed ANOVA "repeated measures designs with replicates", are an extension of the approach in *Biostatistics* Worksheet 300 by adding replicate sets of mached observations. Whereas before, individuals (or objects) comprised "blocks" by themselves, now a block contains replicate sets that are considered more similar to each other (i.e., correlated) than to observations taken from other blocks. The objective of the ANOVA analysis is to control for these correlations when assessing treatment effect. This approach is an example of "mixed-model" (Type III) ANOVA. Shown here is the traditional approach to this problem using a standard two-way ANOVA table and modified Mean Squares F-ratios as specified by Zar's Table 12.3 p. 262 and *Biostatistics* Worksheet 310. More modern approaches to mixed model ANOVA, utilizing liklihood estimation and ANOVA full versus reduced model tests, have largely supplanted the methods presented here. Data is from JC Pinhiero & DM Bates 2004 *Mixed-Effects Models in S and S-Plus* p. 21.

Data Structure:

k treatments exactly matched within individuals (objects). Typically, the order in which specific treatments are presented to individuals is randomized with n replicates for each combination of treatment and block.

Let index i,j indicate the ith column (treatment class) and jth row (block). For each i,j combination, there are n replicates.

	ANOVA	Randomiz	ed Block D	esign				
		Trea	tment Clas					
Blocks	#1	#1 #2 #3		#k				
1	n	n	n		n			
2	n	n	n		n			
3	n	n	n		n			
b	n	n	n		n			
means:	X1bar	X2bar	X3bar		Xkba			

Model:

$$X_{i,i} = \mu + \rho_i + \alpha_i + \alpha_i \rho_i + \varepsilon_{i,i}$$

μ is the grand mean of all objects. ρ_i is a random effect for each block j

 α_i is a constant effect for each treatment i.

 $\alpha_i \rho_i$ is the interaction between treatment and block

 $\boldsymbol{\epsilon}_{i,j}$ is the error term specific to each object i,j

PB Machines example:

 $\sum_{i} \alpha_{i} := 0 \quad \sum_{j} \rho_{j} := 0 \quad \sum_{i} \alpha \rho_{ij} := 0$ ^ allows estimation of parameters.

< where:

Assumptions:

Restriction:

$$\label{eq:rho} \begin{split} \rho_{j} \mbox{ are a random sample} &\sim N(0,\,\sigma_{\rho}^{-2}) \\ \epsilon_{ii} \mbox{ are a random sample} &\sim N(0,\sigma^{2}), \mbox{ spherical} \end{split}$$

 ρ_i and ε_{ii} are independent.

Number & Means:

M := READPRN("c:/DATA/Biostatistics/machines.txt")

$Y := M^{\langle 3 \rangle}$	< measurement variable
k := 3	< number of treatments (machines)
b := 6	< number of blocks (workers)
n := 3	< number of replicates
N := 54	< total number of observations

		1	2	3
	1	1	1	52
	2	1	1	52.8
	3	1	1	53.1
	4	2	1	51.8
	5	2	1	52.8
	6	2	1	53.1
	7	3	1	60
M =	8	3	1	60.2
	9	3	1	58.4
	10	4	1	51.1
	11	4	1	52.3
	12	4	1	50.3
	13	5	1	50.9
	14	5	1	51.8
	15	5	1	51.4
	16	6	1	46.4

$$A_{bar} := \begin{pmatrix} \text{mean}(A_1) \\ \text{mean}(A_2) \\ \text{mean}(A_3) \end{pmatrix} \qquad A_{bar} = \begin{pmatrix} 52,35556 \\ 60,32222 \\ 66,27222 \end{pmatrix} \qquad < \text{means for treatments}$$

$$B_1 := \begin{pmatrix} 52 \\ 52.8 \\ 53.1 \\ 62.1 \\ 62.6 \\ 64 \\ 67.5 \\ 67.2 \\ 66.9 \end{pmatrix} \qquad B_2 := \begin{pmatrix} 51.8 \\ 52.8 \\ 53.1 \\ 59.7 \\ 61.5 \\ 61.5 \\ 61.5 \\ 61.7 \\ 62.3 \end{pmatrix} \qquad B_3 := \begin{pmatrix} 60 \\ 60.2 \\ 58.4 \\ 68.6 \\ 69.7 \\ 70.8 \\ 70.6 \\ 71 \end{pmatrix} \qquad B_4 := \begin{pmatrix} 51.1 \\ 52.3 \\ 50.3 \\ 63.2 \\ 62.8 \\ 62.2 \\ 64.1 \\ 72.1 \\ 72.1 \\ 72.1 \\ 72.1 \\ 72.1 \\ 72.1 \\ 60.2 \\ 61.4 \\ 60.5 \end{pmatrix} \qquad < \text{blocks}$$

$$B_{bar} := \begin{pmatrix} \text{mean}(B_1) \\ \text{mean}(B_2) \\ \text{mean}(B_3) \\ \text{mean}(B_4) \\ \text{mean}(B_5) \\ \text{mean}(B_6) \end{pmatrix} \qquad B_{bar} = \begin{pmatrix} 60.91111 \\ 57.9889 \\ 66.12222 \\ 59.57778 \\ 62.72222 \\ 50.57778 \\ 62.72222 \\ 50.57778 \end{pmatrix} \qquad < \text{means for blocks}$$

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$$AB_{11} := \begin{pmatrix} 52\\ 52.8\\ 53.1 \end{pmatrix} AB_{12} := \begin{pmatrix} 51.8\\ 52.8\\ 53.1 \end{pmatrix} AB_{13} := \begin{pmatrix} 60\\ 60.2\\ 58.4 \end{pmatrix} AB_{14} := \begin{pmatrix} 51.1\\ 52.3\\ 50.3 \end{pmatrix} AB_{15} := \begin{pmatrix} 50.9\\ 51.8\\ 51.4 \end{pmatrix} AB_{16} := \begin{pmatrix} 46.4\\ 44.8\\ 49.2 \end{pmatrix}$$

$$AB_{21} := \begin{pmatrix} 62.1\\ 62.6\\ 64 \end{pmatrix} AB_{22} := \begin{pmatrix} 59.7\\ 60\\ 59 \end{pmatrix} AB_{23} := \begin{pmatrix} 68.6\\ 65.8\\ 69.7 \end{pmatrix} AB_{24} := \begin{pmatrix} 63.2\\ 62.8\\ 62.2 \end{pmatrix} AB_{25} := \begin{pmatrix} 64.8\\ 65\\ 65.4 \end{pmatrix} AB_{26} := \begin{pmatrix} 43.7\\ 44.2\\ 43 \end{pmatrix} < cells$$

$$AB_{31} := \begin{pmatrix} 67.5\\ 67.2\\ 66.9 \end{pmatrix} AB_{32} := \begin{pmatrix} 61.5\\ 61.7\\ 62.3 \end{pmatrix} AB_{33} := \begin{pmatrix} 70.8\\ 70.6\\ 71 \end{pmatrix} AB_{34} := \begin{pmatrix} 64.1\\ 66.2\\ 64 \end{pmatrix} AB_{35} := \begin{pmatrix} 72.1\\ 72\\ 71.1 \end{pmatrix} AB_{36} := \begin{pmatrix} 62\\ 61.4\\ 60.5 \end{pmatrix}$$

$$(mean(AB_{11}) mean(AB_{21}) mean(AB_{31})) mean(AB_{32}) mean(AB_{33}) mean(AB_{33}) = AB_{4} = \begin{pmatrix} 52.6333 & 62.9 & 67.2\\ 52.5667 & 59.5667 & 61.8333\\ 59.5333 & 68.0333 & 70.8 \end{pmatrix} < cell means$$

$$AB := \begin{pmatrix} mean(AB_{13}) & mean(AB_{23}) & mean(AB_{33}) \\ mean(AB_{14}) & mean(AB_{24}) & mean(AB_{34}) \\ mean(AB_{15}) & mean(AB_{25}) & mean(AB_{35}) \\ mean(AB_{16}) & mean(AB_{26}) & mean(AB_{36}) \end{pmatrix} AB = \begin{pmatrix} 59.5353 & 68.0333 & 70.8 \\ 51.2333 & 62.7333 & 64.7667 \\ 51.3667 & 65.0667 & 71.7333 \\ 46.8 & 43.6333 & 61.3 \end{pmatrix} < cell mean \\$$

Sums of Squares:

SSA for Treatment:

i := 1..k $SSA := n \cdot b \cdot \sum_{i} (A_{bar_i} - GM)^2$ SSA = 1755.2633

SSB for factor B:

j := 1 .. b

 $SSB := n \cdot k \cdot \sum_{j} \left(B_{bar_{j}} - GM \right)^{2} \qquad SSB = 1241.895$

SSAB for Interaction AB:

$$SSAB := n \cdot \sum_{i} \sum_{j} \left(AB_{j,i} - A_{bar_{i}} - B_{bar_{j}} + GM \right)^{2} \qquad SSAB = 426.53$$

SS Total:

kk := 1 .. N

SSTO :=
$$\sum_{kk} (Y_{kk} - GM)^2$$
 SSTO = 3456.975

SSE for Error:

SSE := SSTO - SSAB - SSA - SSB SSE = 33.2867

Randomized Block NOVA Table:

Sum of Squares:	Degrees of Fre	edom:	Mean Squares:		
SSA = 1755.2633	$df_A := k - 1$	$df_A = 2$	$MSA := \frac{SSA}{df_A}$	MSA = 877.6317	
SSB = 1241.895	$df_B := (b - 1)$	$df_B = 5$	$MSB := \frac{SSB}{df_B}$	MSB = 248.379	
SSAB = 426.53	$df_{AB} \coloneqq (k-1) \cdot (b-1)$	$df_{AB} = 10$	$MSAB := \frac{SSAB}{df_{AB}}$	MSAB = 42.653	
SSE = 33.2867	$df_E \coloneqq k \cdot b \cdot (n-1)$	$df_E = 36$	$MSE := \frac{SSE}{df_E}$	MSE = 0.9246	
SSTO = 3456.975	$df_T \coloneqq n \cdot k \cdot b - 1$	$df_{T} = 53$			

Omnibus F Test for Treatment Effect:

Hypotheses:

$H_0: \alpha_i = 0$ for all i	< All treatment class deviations from the grand mean are 0
\mathbf{H}_{1} : At least one $\alpha_{i} \neq 0$	< Two sided test

Test Statistic:

 $F := \frac{MSA}{MSAB}$ F = 20.5761
 Ratio of "treatment" versus "interaction" Mean Squares

Sampling Distribution of the test Statistic F:

If Assumptions hold and H_0 is true then F $\sim F_{(dfA,dfAB)}$

Critical Value of the Test:

 $\alpha := 0.05$

 $C := qF(1 - \alpha, df_A, df_{AB}) \qquad \qquad C = 4.1028$

Decision Rule:

IF F > C, THEN REJECT H_0 OTHERWISE ACCEPT H_0

Probability Value:

 $P_A := 1 - pF(F, df_A, df_{AB})$ $P_A = 0.0002855$

Omnibus F Test for Block Effect:

Hypotheses:

$H_0: \rho_j = 0$ for all j	< All treatment class deviations from the grand mean are 0
$\mathbf{H}_{1}: At \ least \ one \ \rho_{j} \neq 0$	< Two sided test

Test Statistic:

$F := \frac{MSB}{MSE}$	F = 268.6254	< Ratio of "block" versus "error" Mean Squares
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Sampling Distribution of the test Statistic F:

If Assumptions hold and H_0 is true then F ~F_(dfB,dfE)

Critical Value of the Test:

 $\alpha := 0.05$ < Probability of Type I error must be explicitly set

 $C := qF(1 - \alpha, df_B, df_E) \qquad C = 2.4772$

Decision Rule:

IF F > C, THEN REJECT H_0 OTHERWISE ACCEPT H_0

Probability Value:

 $P_A := 1 - pF(F, df_B, df_E) \qquad P_A = 0$

F Test for Treatment X Block Interaction:

Hypotheses:

$H_0: \alpha \rho_{ij} = 0$ for all ij	< All treatment class deviations from the grand mean are 0
$\mathbf{H}_{\mathbf{i}}: At \ least \ one \ \alpha \rho_{\mathbf{ij}} \neq 0$	< Two sided test

Test Statistic:

$F := \frac{MSAB}{MSE}$	F = 46.1298	< Ratio of "interaction" versus "error" Mean Squares
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Sampling Distribution of the test Statistic F:

If Assumptions hold and H_0 is true then F ~F_(dfAB,dfE)

Critical Value of the Test:

 $\alpha := 0.05$ < Probability of Type I error must be explicitly set

 $C := qF(1 - \alpha, df_{AB}, df_E) \qquad C = 2.1061$

Decision Rule:

IF F > C, THEN REJECT H_0 OTHERWISE ACCEPT H_0

Probability Value:

 $P_A := 1 - pF(F, df_{AB}, df_E) \qquad P_A = 0$

Prototype in R:

#ANOVA FOR RANDOMIZED BLOCKS WITH REPLICATES

M=read.table("C:/DATA/Biostatistics/machinesR.txt",header=TRUE) M attach(M) fMachine=factor(Machine) fWorker=factor(Worker)

LM=lm(score~fMachine*fWorker) anova(LM)

> anova(LM)

alpha = 0.05	Analysis of Vari	ance	a Tabl	.e					
MSA=anova(LM)[1,3] MSB=anova(LM)[2,3]	Response: score								
MSAB=anova(LM)[3,3]		Df	Sum	$\mathbf{S}\mathbf{q}$	Mean Sq	F value		Pr (>F)	
	fMachine	2	1755.	26	877.63	949.17	<	2.2e-16	***
MSE=anova(LM)[4,3]	fWorker	5	1241.	89	248.38	268.63	<	2.2e-16	***
dfA=anova(LM)[1,1]	fMachine:fWorker	10	426.	53	42.65	46.13	<	2.2e-16	***
dfB=anova(LM)[2,1]	Residuals	36	33.	29	0.92				
dfAB=anova(LM)[3,1]									
dfE=anova(LM)[4,1]	Signif. codes:	0 ';	***′0	0.0	01 `**'	0.01 `*'	0	.05 `.′ (0.1 ` ′ 1

> F

> C

> P

> F

> C

> P

> F

> C

> P

[1] 0

[1] 0

[1] 20.57608

[1] 4.102821

[1] 268.6254

[1] 2.477169

[1] 46.12982

[1] 2.106054

[1] 0.0002855485

#OMNIBUS F TEST FOR TREATMENT EFFECT:
F=MSA/MSAB
F
C=qf(1-alpha,dfA,dfAB)
C
P=1-pf(F,dfA,dfAB)
Ρ

#OMNIBUS F TEST FOR BLOCK EFFECT: F=MSB/MSE F C=qf(1-alpha,dfB,dfE)

C P=1-pf(F,dfB,dfE) P

#F TEST FOR TREATMENT BY BLOCK INTERACTION: F=MSAB/MSE F C=qf(1-alpha,dfAB,dfE) C P=1-pf(F,dfAB,dfE)

Ρ