

ORIGIN = 1

ANOVA F-test/t-test for Simple Linear Regression and Interval Estimation

Goodness of fit of a fitted regression line can be tested using the F-test for Regression (also known as the ANOVA for Regression). Alternatively, and equivalently, fit of the regression can be tested using a t-test approach. The latter method also allows estimation of confidence intervals for the slope parameter β . Interval estimates can also be derived for the Regression line itself as a *mean*, as well as for *prediction* of "new" observations.

```
ZAR := READPRN("c:/DATA/Biostatistics/ZarEX17.1R.txt")
```

```
X := ZAR(2)
```

^Zar Example 17.1

```
Y := ZAR(3)
```

```
Xbar := mean(X) Xbar = 10
```

```
Ybar := mean(Y) Ybar = 3.4154
```

```
s := sqrt(Var(Y)) s = 1.2799 s2 = 1.6381
```

```
n := length(Y) n = 13
```

```
i := 1 .. n
```

$$X = \begin{pmatrix} 3 \\ 4 \\ 5 \\ 6 \\ 8 \\ 9 \\ 10 \\ 11 \\ 12 \\ 14 \\ 15 \\ 16 \\ 17 \end{pmatrix} \quad Y = \begin{pmatrix} 1.4 \\ 1.5 \\ 2.2 \\ 2.4 \\ 3.1 \\ 3.2 \\ 3.2 \\ 3.9 \\ 4.1 \\ 4.7 \\ 4.5 \\ 5.2 \\ 5 \end{pmatrix}$$

	1	2	3
1	1	3	1.4
2	2	4	1.5
3	3	5	2.2
4	4	6	2.4
5	5	8	3.1
6	6	9	3.2
7	7	10	3.2
8	8	11	3.9
9	9	12	4.1
10	10	14	4.7
11	11	15	4.5
12	12	16	5.2
13	13	17	5

Assumptions:

- Standard Linear Regression depends on specifying in advance which variable is to be considered 'dependent' and which 'independent'. This decision matters as changing roles for Y & X usually produces a different result.
- $Y_1, Y_2, Y_3, \dots, Y_n$ (dependent variable) is a random sample $\sim N(\mu, \sigma^2)$.
- $X_1, X_2, X_3, \dots, X_n$ (independent variable) with each value of X_i matched to Y_i

Model:

where: α is the y **intercept** of the regression line (translation)

β is the **slope** of the regression line (scaling coefficient)

ϵ is the error factor in prediction of Y given that it is a random variable with $N(0, \sigma^2)$

Corrected Sums of Squares:

$$L_{xx} := \sum_i (X_i - X_{\bar{}})^2$$

$$L_{yy} := \sum_i (Y_i - Y_{\bar{}})^2$$

$$L_{xy} := \sum_i (X_i - X_{\bar{}}) \cdot (Y_i - Y_{\bar{}})$$

$$L_{xx} = 262$$

$$L_{yy} = 19.6569$$

$$L_{xy} = 70.8$$

Estimated Regression Coefficients for $Y = \alpha + \beta X$:

$$b := \frac{L_{xy}}{L_{xx}} \quad b = 0.2702 \quad < \text{sample estimate of } \beta$$

$$a := Y_{\bar{}} - b \cdot X_{\bar{}} \quad a = 0.7131 \quad < \text{sample estimate of } \alpha$$

Estimated values of Y (\hat{Y}_i):

$$\hat{Y}_i := a + b \cdot X_i$$

Residuals:

$$e_i := \hat{Y}_i - Y_i$$

Sum of Squares Decomposition of the Regression:

Once a regression model ($Y = \alpha + \beta X + \varepsilon$) is fitted with data as $\hat{Y}_i = a + bX_i$, one still needs to determine how useful the regression might be, especially whether knowledge about the X_i provide insight into interpreting the Y_i as a random variable from a Normal distribution with error ε_i .

This is done by considering a "partition" of total Sums of Squares in the sample of Y_i .

$$SST := \sum_i (Y_i - \bar{Y})^2 \quad SST = 19.6569 \quad L_{yy} = 19.6569 \quad < \text{Total Sum of Squares}$$

$$SSR := \sum_i (\hat{Y}_i - \bar{Y})^2 \quad SSR = 19.1322 \quad < \text{Regression Sum of Squares}$$

$$SSE := \sum_i (Y_i - \hat{Y}_i)^2 \quad SSE = 0.5247 \quad < \text{Residual (also called "Error") Sum of Squares}$$

These Sums of Squares tally as follows:

$$SST = 19.6569 \quad SSR + SSE = 19.6569$$

And the ratio of SSR to SSE can be used as a measure of "fit" of the data to the regression.

ANOVA Table for Linear Regression:

	SS	df	MS	
Regression:	$SSR = 19.1322$	$df_R := 1$	$MSR := \frac{SSR}{df_R}$	$MSR = 19.1322$
Residual:	$SSE = 0.5247$	$df_E := n - 2$	$MSE := \frac{SSE}{df_E}$	$MSE = 0.0477$
TOTAL:	$SST = 19.6569$	$df_T := n - 1$	$MST := \frac{SST}{df_T}$	$MST = 1.6381$

ANOVA F-test for Regression Slope:

Hypotheses:

$$\begin{aligned} H_0: \beta = 0 & \quad < \text{Slope of the Regression is zero implying no relationship between } X_i \text{ and } Y_i \\ H_1: \beta \neq 0 & \quad \text{Two sided test} \end{aligned}$$

Test Statistic:

$$F := \frac{MSR}{MSE} \quad F = 401.0875 \quad < F \text{ is the ratio of sample variances}$$

Sampling Distribution:

If Assumptions hold and H_0 is true, then $F_s \sim F_{(1,n-2)}$

Critical Value of the Test:

$$\alpha := 0.05 \quad < \text{Probability of Type I error must be explicitly set}$$

$$C := q_F(1 - \alpha, 1, n - 2) \quad C = 4.8443$$

Decision Rule:

IF $F > C$, THEN REJECT H_0 , OTHERWISE ACCEPT H_0

Probability Value:

$$P := 1 - pF(F, 1, n - 2) \quad P = 5.2671 \times 10^{-10}$$

Prototype in R:

```
#SIMPLE LINEAR REGRESSION
#ZAR EXAMPLE 17.1
ZAR=read.table("c:/DATA/Biostatistics/ZarEX17.1R.txt")
ZAR
attach(ZAR)
X=ageX
Y=wingLY

#CREATING LINEAR MODEL LM:
LM=lm(Y~X)
Yhat=fitted(LM)
e=residuals(LM)
RESULTS=data.frame(X,Y,Yhat,e)
RESULTS

#ANOVA F-TEST:
anova(LM)
```

Analysis of Variance Table					
		Response: Y			
	Df	Sum Sq	Mean Sq	F value	Pr(>F)
X	1	19.1322	19.1322	401.09	5.267e-10 ***
Residuals	11	0.5247	0.0477		

				Signif. codes:	0 '****' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

The t-Test Approach and Interval Estimation for Simple Linear Regression:

The t-test approach is exactly equivalent to the F-test above with the statistic $F = \text{statistic } t^2$. In contrast to the F-test, however, the t-test allows for one-tailed tests and for construction of confidence intervals of the regression (to capture β), confidence intervals for the regression estimates (to capture \hat{Y}_i) and prediction intervals for new observations (to capture new observations X_n).

t-Test for Regression Parameter β (slope):**Hypotheses:**

- $H_0: \beta = 0$ < Slope of the Regression is zero implying no relationship between X_i and Y_i
- $H_1: \beta \neq 0$ < Two Sided Case
- $H_1: \beta < 0$ < One Sided Case Lower Tail)
- $H_1: \beta > 0$ < One Sided Case (Upper Tail)

Test Statistic:

$$t_\beta := \frac{b}{\sqrt{\frac{MSE}{L_{xx}}}} \quad t_\beta = 20.0272$$

< b is unbiased point estimate of β
< MSE is Mean Square Error from ANOVA table also denoted $s^2_{x,y}$
< L_{xx} Corrected sums of squares of X as defined in Regression

Sampling Distribution of Test Statistic t_β :

If Assumptions hold and H_0 is true, then $t_\beta \sim t_{(n-2)}$

Critical Values of the Test:

$$\alpha := 0.05 \quad < \text{Probability of Type I error must be explicitly set}$$

$$C_1 := qt\left(\frac{\alpha}{2}, n - 2\right) \quad C_1 = -2.201 \quad < \text{Two sided lower Critical Value}$$

$$C_2 := qt\left(1 - \frac{\alpha}{2}, n - 2\right) \quad C_2 = 2.201 \quad < \text{Two sided upper Critical Value}$$

$$C := |C_2| \quad C = 2.201 \quad < \text{Critical value used for two sided test (to simplify)}$$

$$C_3 := qt(\alpha, n - 2) \quad C_3 = -1.7959 \quad < \text{One sided lower Critical Value}$$

$$C_4 := qt(1 - \alpha, n - 2) \quad C_4 = 1.7959 \quad < \text{One sided upper Critical Value}$$

Decision Rules:

IF $|t_\beta| > C$ THEN REJECT H_0 , OTHERWISE ACCEPT H_0 < Two sided case

IF $t_\beta < C_3$, THEN REJECT H_0 , OTHERWISE ACCEPT H_0 < One sided case lower tail

IF $t_\beta > C_4$, THEN REJECT H_0 , OTHERWISE ACCEPT H_0 < One sided case upper tail

Probability Values:

$$P_\beta := 2 \cdot pt(t_\beta, n - 2) \quad P_\beta = 2 \quad < \text{if } t \leq 0$$

$$P_\beta := 2 \cdot (1 - pt(t_\beta, n - 2)) \quad P_\beta = 5.267053 \times 10^{-10} \quad < \text{if } t > 0 \quad < \text{Two sided case}$$

$$P_\beta := pt(t_\beta, n - 2) \quad P_\beta = 1 \quad < \text{One sided case lower tail}$$

$$P_\beta := 1 - pt(t_\beta, n - 2) \quad P_\beta = 2.633527 \times 10^{-10} \quad < \text{One sided case upper tail}$$

Confidence Interval for the Regression Slope (β):

$$CI := \left(b - C \cdot \sqrt{\frac{MSE}{L_{xx}}} \quad b + C \cdot \sqrt{\frac{MSE}{L_{xx}}} \right) \quad CI = (0.2405308 \quad 0.2999272) \quad < \text{Two sided case} \quad b = 0.2702$$

$$CIL := b - C_3 \cdot \sqrt{\frac{MSE}{L_{xx}}} \quad \text{minus infinity to} \quad CIL = 0.294461 \quad < \text{One sided case lower tail}$$

$$CIU := b - C_4 \cdot \sqrt{\frac{MSE}{L_{xx}}} \quad CIU = 0.245997 \quad \text{to infinity} \quad < \text{One sided case upper tail}$$

t-Test for Regression Parameter α (intercept):

Hypotheses:

- $\alpha_0 := 0$ < any value may be tested
 $H_0: \alpha = \alpha_0$ < tests for intercept $= \alpha_0 = 0$ here, but may be set to other values as desired
 $H_1: \alpha \neq \alpha_0$ < Two Sided Case
 $H_1: \alpha_0 < 0$ < One Sided Case Lower Tail)
 $H_1: \alpha_0 > 0$ < One Sided Case (Upper Tail)

Test Statistic:

$$t_\alpha := \frac{a - \alpha_0}{\sqrt{MSE \cdot \left(\frac{1}{n} + \frac{\bar{X}_\text{bar}^2}{L_{XX}} \right)}} \quad t_\alpha = 4.8213$$

Sampling Distribution of Test Statistic t_α :

If Assumptions hold and H_0 is true, then $t_\alpha \sim t_{(n-2)}$

Critical Values of the Test:

$$\alpha := 0.05 \quad < \text{Probability of Type I error must be explicitly set}$$

$$C_1 := qt\left(\frac{\alpha}{2}, n-2\right) \quad C_1 = -2.201$$

$$C_2 := qt\left(1 - \frac{\alpha}{2}, n-2\right) \quad C_2 = 2.201 \quad < \text{Note degrees of freedom} = (n-2)$$

$$C := |C_1| \quad C = 2.201$$

Decision Rules:

- | | |
|--|-----------------------------|
| IF $ t_\alpha > C$ THEN REJECT H_0 , OTHERWISE ACCEPT H_0 | < Two sided case |
| IF $t_\alpha < C_3$, THEN REJECT H_0 , OTHERWISE ACCEPT H_0 | < One sided case lower tail |
| IF $t_\alpha > C_4$, THEN REJECT H_0 , OTHERWISE ACCEPT H_0 | < One sided case upper tail |

Probability Values:

$P_\alpha := 2 \cdot pt(t_\alpha, n-2)$	$P_\alpha = 1.9995$	< if $t \leq 0$
$P_\alpha := 2 \cdot (1 - pt(t_\alpha, n-2))$	$P_\alpha = 0.000535$	< if $t > 0$ < Two sided case
$P_\alpha := pt(t_\alpha, n-2)$	$P_\alpha = 0.999733$	< One sided case lower tail
$P_\alpha := 1 - pt(t_\alpha, n-2)$	$P = 5.267053 \times 10^{-10}$	< One sided case upper tail

Confidence Interval for Intercept (α):

$$a = 0.7131$$

$$CI := \left[a - C \cdot \sqrt{MSE \cdot \left(\frac{1}{n} + \frac{\bar{X}^2}{L_{XX}} \right)}, a + C \cdot \sqrt{MSE \cdot \left(\frac{1}{n} + \frac{\bar{X}^2}{L_{XX}} \right)} \right]$$

$$CI = (0.387559 \quad 1.03863) \quad < \text{Two sided case}$$

$$CIL := a - C_3 \cdot \sqrt{MSE \cdot \left(\frac{1}{n} + \frac{\bar{X}^2}{L_{XX}} \right)}$$

minus infinity to CIL = 0.978714 < One sided case lower tail

$$CIU := a - C_4 \cdot \sqrt{MSE \cdot \left(\frac{1}{n} + \frac{\bar{X}^2}{L_{XX}} \right)}$$

CIU = 0.447475 to infinity < One sided case upper tail

Prototype in R:

#t-TESTS FOR SLOPE & INTERCEPT:
summary(LM)

$$F = 401.0875$$

$$t_\beta = 20.0272 \quad t_\alpha = 4.8213$$

$$t_\beta^2 = 401.0875$$

[^] Note for test of β , (slope):

$$\text{Statistics } F = t_\beta^2$$

#CONFIDENCE INTERVAL FOR SLOPE:

#CALCULATED BY HAND!

n=length(Y)

Lxx=sum((X-mean(X))^2)

Lxx

Lxy=sum((X-mean(X))*(Y-mean(Y)))

Lxy

b=Lxy/Lxx

b

MSE=summary(LM)\$sigma^2

MSE

alpha=0.05

C=abs(qt(1-alpha/2,n-2))

C

stderr=sqrt(MSE/Lxx)

CIL=b-C*stderr

CIL

CIU=b+C*stderr

CIU

CONFIDENCE INTERVALS FOR

#REGRESSION PARAMETERS

#THE EASIER WAY:

confint(lm(Y~X),level=0.95)

> summary(LM)

Call:

lm(formula = Y ~ X)

Residuals:

Min	1Q	Median	3Q	Max
-0.30699	-0.21538	0.06553	0.16324	0.22507

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.71309	0.14790	4.821	0.000535 ***
X	0.27023	0.01349	20.027	5.27e-10 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.2184 on 11 degrees of freedom

Multiple R-squared: 0.9733, Adjusted R-squared: 0.9709

F-statistic: 401.1 on 1 and 11 DF, p-value: 5.267e-10

> Lxx

[1] 262

> Lxy

[1] 70.8

> MSE

[1] 0.04770085

> CIL

[1] 0.2405308

> CIU

[1] 0.2999272

> confint(lm(Y~X),level=0.95)

	2.5 %	97.5 %
(Intercept)	0.3875590	1.0386300
X	0.2405308	0.2999272

Confidence Interval for Regression Estimates \hat{Y}_i and New Predictions of Y_i :

One or more values of X_i must be explicitly specified to obtain a prediction CI for \hat{Y}_i :

$$i := 1..n$$

$X_{ni} := X_i$ < here using all original values of X , but any X values may be specified instead...

Confidence Interval for Regression (CI):

$$CI := \left[\hat{Y}_i - C \cdot \sqrt{MSE \cdot \left[\frac{1}{n} + \frac{(X_n - \bar{X})^2}{L_{xx}} \right]} \quad \hat{Y}_i + C \cdot \sqrt{MSE \cdot \left[\frac{1}{n} + \frac{(X_n - \bar{X})^2}{L_{xx}} \right]} \right] \quad < \text{Two sided case}$$

$$CIL := \hat{Y}_i - C_3 \cdot \sqrt{MSE \cdot \left[\frac{1}{n} + \frac{(X_n - \bar{X})^2}{L_{xx}} \right]} \quad < \text{One sided case lower tail}$$

$$CIU := \hat{Y}_i + C_4 \cdot \sqrt{MSE \cdot \left[\frac{1}{n} + \frac{(X_n - \bar{X})^2}{L_{xx}} \right]} \quad < \text{One sided case upper tail}$$

minus infinity to

$\hat{Y}_i =$	1.5238	1.2768	1.7707	1.725293	1.32227
	1.794	1.5715	2.0166	1.975596	1.612425
	2.0642	1.8647	2.2638	2.227071	1.901408
	2.3345	2.1559	2.513	2.480171	2.188766
	2.8749	2.729	3.0209	2.994019	2.755834
	3.1452	3.0086	3.2817	3.256607	3.033704
	3.4154	3.2821	3.5487	3.52417	3.306599
	3.6856	3.549	3.8222	3.797065	3.574162
	3.9558	3.8099	4.1018	4.074935	3.83675
	4.4963	4.3177	4.6749	4.642004	4.350598
	4.7665	4.567	4.9661	4.929361	4.603698
	5.0368	4.8142	5.2593	5.218344	4.855173
	5.307	5.06	5.554	5.508499	5.105477

to infinity

Prototype in R:

#CONFIDENCE INTERVALS FOR \hat{Y}_i :

CONF=predict(lm(Y~X),interval="confidence",level=0.95)

CN=data.frame(CONF)

CN

> CN

	fit	lwr	upr
1	1.523782	1.276815	1.770748
2	1.794011	1.571465	2.016556
3	2.064240	1.864678	2.263801
4	2.334469	2.155899	2.513038
5	2.874927	2.728970	3.020883
6	3.145156	3.008564	3.281747
7	3.415385	3.282061	3.548709
8	3.685614	3.549022	3.822205
9	3.955843	3.809886	4.101799
10	4.496301	4.317731	4.674870
11	4.766530	4.566968	4.966091
12	5.036759	4.814213	5.259304
13	5.306988	5.060021	5.553954

Prediction Interval for Regression (PI):

$$PI := \left[Y_{\text{hat}} - C \cdot \sqrt{MSE \cdot \left[1 + \frac{1}{n} + \frac{(X_n - X_{\bar{}})^2}{L_{XX}} \right]} \quad Y_{\text{hat}} + C \cdot \sqrt{MSE \cdot \left[1 + \frac{1}{n} + \frac{(X_n - X_{\bar{}})^2}{L_{XX}} \right]} \right] \quad < \text{Two sided case}$$

$$PIL := Y_{\text{hat}} - C_3 \cdot \sqrt{MSE \cdot \left[1 + \frac{1}{n} + \frac{(X_n - X_{\bar{}})^2}{L_{XX}} \right]} \quad < \text{One sided case lower tail}$$

$$PIU := Y_{\text{hat}} - C_4 \cdot \sqrt{MSE \cdot \left[1 + \frac{1}{n} + \frac{(X_n - X_{\bar{}})^2}{L_{XX}} \right]} \quad < \text{One sided case upper tail}$$

minus infinity to

$Y_{\text{hat}} =$	$\begin{pmatrix} 1.5238 \\ 1.794 \\ 2.0642 \\ 2.3345 \\ 2.8749 \\ 3.1452 \\ 3.4154 \\ 3.6856 \\ 3.9558 \\ 4.4963 \\ 4.7665 \\ 5.0368 \\ 5.307 \end{pmatrix}$	$PI =$	$\begin{pmatrix} 0.9833 & 2.0642 \\ 1.2643 & 2.3237 \\ 1.5438 & 2.5847 \\ 1.8217 & 2.8473 \\ 2.3726 & 3.3773 \\ 2.6454 & 3.6449 \\ 2.9165 & 3.9142 \\ 3.1859 & 4.1853 \\ 3.4535 & 4.4582 \\ 3.9835 & 5.0091 \\ 4.246 & 5.287 \\ 4.507 & 5.5665 \\ 4.7666 & 5.8474 \end{pmatrix}$	$PIL =$	$\begin{pmatrix} 1.964748 \\ 2.226235 \\ 2.488926 \\ 2.752887 \\ 3.284839 \\ 3.552913 \\ 3.822422 \\ 4.093371 \\ 4.365755 \\ 4.914719 \\ 5.191217 \\ 5.468983 \\ 5.747954 \end{pmatrix}$	$PIU =$	$\begin{pmatrix} 1.082815 \\ 1.361786 \\ 1.639553 \\ 1.91605 \\ 2.465015 \\ 2.737398 \\ 3.008348 \\ 3.277856 \\ 3.545931 \\ 4.077882 \\ 4.341843 \\ 4.604534 \\ 4.866021 \end{pmatrix}$
--------------------	--	--------	--	---------	--	---------	---

to infinity

Prototype in R:

#PREDICTION INTERVALS FOR X_n :

PRED=predict(lm(Y~X),interval="prediction",level=0.95)

PR=data.frame(PRED)

PR

> PR

	fit	lwr	upr
1	1.523782	0.9833454	2.064218
2	1.794011	1.2642884	2.323733
3	2.064240	1.5437554	2.584724
4	2.334469	1.8216666	2.847271
5	2.874927	2.3725501	3.377303
6	3.145156	2.6454195	3.644892
7	3.415385	2.9165317	3.914238
8	3.685614	3.1858775	4.185350
9	3.955843	3.4534661	4.458219
10	4.496301	3.9834986	5.009103
11	4.766530	4.2460455	5.287014
12	5.036759	4.5070365	5.566481
13	5.306988	4.7665515	5.847424

> PRED=predict(lm(Y~X),interval="prediction",level=0.95)

Warning message:

In predict.lm(lm(Y ~ X), interval = "prediction", level = 0.95) :

Predictions on current data refer to _future_ responses

#PLOTTING INTERVALS IN R:

```
plot(X,Y)
abline(lm(Y~X),col="blue")
segments(X,PR$lwr,X,PR$upr,col="red")
segments(X,CN$lwr,X,CN$upr,col="green")
points(X,CN$lwr,col="green")
points(X,CN$upr,col="green")
points(X,PR$lwr,col="red")
points(X,PR$upr,col="red")
points(X,PR$fit,col="blue")
```

