

## Nested Two-way ANOVA Model - Balanced Case

Nested ANOVA designs occur when, either by efficiency of design or by necessity, not all combinations of factors are studied. Nested design in a two-way ANOVA (involving two factors) permits analysis of the independent factor (factor A below) in the same way as a crossed two-way ANOVA. The difference lies in factor B nested within A. Here effects of different cell blocks within single treatment levels of A are assessed much as if in single factor ANOVA. Example below comes from Chapter 26 in Kuter et al. (KNNL) *Applied Linear Statistical Models* 5th Edition. See particularly Table 26.2 for discussion of how this design relates to one-way (single-factor) ANOVA.

### Data Structure:

Data are structured such that different levels of Nested Factor B occur only within single levels of independent Factor A. Other possible combinations are not studied.

Let index  $i, j$  indicate the  $i$ th row (treatment classes of Variable A) and  $j$ th column (treatment classes of Variable B)

Treatment Classes of Variable A:	Nested Two-Way ANOVA					
	Treatment Classes of Variable B Nested Within A:					
	#1	#2	#3	#4	#5	#6
#1	n	n				
#2			n	n		
#3					n	n
Each cell consists of n replicates with means $\bar{Y}_{ij}$						

Also let:  $\bar{Y}_{i\cdot}$  = mean over all columns for  $A_i$ .  
 $\bar{Y}_{ij}$  = mean within data blocks.  
 $\bar{Y}_{\cdot\cdot}$  = overall mean.

### Model:

$$Y_{ijk} = \mu_{\cdot\cdot} + \alpha_i + \beta_{j(i)} + \epsilon_{ijk}$$

where:

Note: this model has no associated test for interactions because of the nested design.

### Restrictions:

$$\sum_i \alpha_i := 0 \quad \sum_j \beta_{j \text{ in } i} := 0$$

$\mu_{\cdot\cdot}$  is a constant = grand mean of all objects.  
 $\alpha_i$  are effect coefficients for classes  $i$  in Variable A.  
 $\beta_j$  are nested coefficients for classes  $j$  of Variable B in A.  
 $\epsilon_{ijk}$  is the error term specific to each object  $i, j, k$   
 $i = 1$  to  $a$ ,  $j = 1$  to  $b$ ,  $k = 1$  to  $n$ ,  $N = abn$

### Assumptions:

- $\epsilon_{ijk}$  are a random sample  $\sim N(0, \sigma^2)$
- variance is homogeneous across cells

KNNL dataset Table 26.1

### Example:

```
K := READPRN("c:/2008LinearModelsData/TrainingSchoolR.txt")
```

$$K = \begin{pmatrix} 1 & 25 & 1 & 1 \\ 2 & 29 & 1 & 1 \\ 3 & 14 & 1 & 2 \\ 4 & 11 & 1 & 2 \\ 5 & 11 & 2 & 1 \\ 6 & 6 & 2 & 1 \\ 7 & 22 & 2 & 2 \\ 8 & 18 & 2 & 2 \\ 9 & 17 & 3 & 1 \\ 10 & 20 & 3 & 1 \\ 11 & 5 & 3 & 2 \\ 12 & 2 & 3 & 2 \end{pmatrix}$$

### Variable Assignment:

$$Y := K^{(2)} \quad \begin{matrix} < \text{response (dependent) variable } Y \\ B_1 & B_2 \end{matrix}$$

$$A_1 \quad Y_{11} := \begin{pmatrix} 25 \\ 29 \end{pmatrix} \quad Y_{12} := \begin{pmatrix} 14 \\ 11 \end{pmatrix}$$

$$A_2 \quad Y_{21} := \begin{pmatrix} 11 \\ 6 \end{pmatrix} \quad Y_{22} := \begin{pmatrix} 22 \\ 18 \end{pmatrix}$$

$$A_3 \quad Y_{31} := \begin{pmatrix} 17 \\ 20 \end{pmatrix} \quad Y_{32} := \begin{pmatrix} 5 \\ 2 \end{pmatrix}$$

< response blocks

$$Y_{\text{blocks}} := \begin{pmatrix} Y_{11} & Y_{12} \\ Y_{21} & Y_{22} \\ Y_{31} & Y_{32} \end{pmatrix}$$

$$a := 3 \quad b := 2 \quad n := 2 \quad N := \text{length}(Y)$$

**Means:**

$$\mu_{\text{dotdot}} := \text{mean}(Y)$$

$$\mu_{\text{dotdot}} = 15 \quad < \text{grand mean}$$

$$\mu := \begin{pmatrix} \text{mean}(Y_{11}) & \text{mean}(Y_{12}) \\ \text{mean}(Y_{21}) & \text{mean}(Y_{22}) \\ \text{mean}(Y_{31}) & \text{mean}(Y_{32}) \end{pmatrix}$$

$$\mu = \begin{pmatrix} 27 & 12.5 \\ 8.5 & 20 \\ 18.5 & 3.5 \end{pmatrix} \quad < \text{block means}$$

$$\mu_{\text{idot}} := \begin{pmatrix} \text{mean}(\mu_{1,1}, \mu_{1,2}) \\ \text{mean}(\mu_{2,1}, \mu_{2,2}) \\ \text{mean}(\mu_{3,1}, \mu_{3,2}) \end{pmatrix}$$

$$\mu_{\text{idot}} = \begin{pmatrix} 19.75 \\ 14.25 \\ 11 \end{pmatrix} \quad < \text{Independent Factor A means (row means)}$$

**Model Parameters:**

$$\alpha := \mu_{\text{idot}} - \mu_{\text{dotdot}}$$

$$\alpha = \begin{pmatrix} 4.75 \\ -0.75 \\ -4 \end{pmatrix} \quad < \alpha \text{ coefficients for Independent Factor A}$$

$$\beta_{\text{BinA}} := \begin{pmatrix} \mu_{1,1} - \mu_{\text{idot}_1} & \mu_{1,2} - \mu_{\text{idot}_1} \\ \mu_{2,1} - \mu_{\text{idot}_2} & \mu_{2,2} - \mu_{\text{idot}_2} \\ \mu_{3,1} - \mu_{\text{idot}_3} & \mu_{3,2} - \mu_{\text{idot}_3} \end{pmatrix}$$

$$\beta_{\text{BinA}} = \begin{pmatrix} 7.25 & -7.25 \\ -5.75 & 5.75 \\ 7.5 & -7.5 \end{pmatrix} \quad < \beta \text{ coefficients for Nested Factor B in A}$$

**Fitted Values & Residuals:**

$$Y = \begin{pmatrix} 25 \\ 29 \\ 14 \\ 11 \\ 11 \\ 6 \\ 22 \\ 18 \\ 17 \\ 20 \\ 5 \\ 2 \end{pmatrix} \quad \mu_{\text{vec}} := \begin{pmatrix} 27 \\ 27 \\ 12.5 \\ 12.5 \\ 8.5 \\ 8.5 \\ 20 \\ 20 \\ 18.5 \\ 18.5 \\ 3.5 \\ 3.5 \end{pmatrix}$$

$$e := Y - \mu_{\text{vec}} \quad e = \begin{pmatrix} -2 \\ 2 \\ 1.5 \\ -1.5 \\ 2.5 \\ -2.5 \\ 2 \\ -2 \\ -1.5 \\ 1.5 \\ 1.5 \\ -1.5 \end{pmatrix} \quad < e = \text{residuals}$$

**Sums of Squares:**

^ means for each block in vector form associated with each value of Y

**Σ form:**

$$SSA := b \cdot n \cdot \sum_{i=1}^a (\alpha_i)^2$$

$$SSA = 156.5$$

**Vector form:**

$$b \cdot n \cdot (\alpha^T \cdot \alpha) = (156.5)$$

$$SSB_{\text{inA}} := n \cdot \sum_{i=1}^a \sum_{j=1}^b (\beta_{\text{BinA}_{i,j}})^2$$

$$SSB_{\text{inA}} = 567.5$$

$$\beta := \text{stack}(\beta_{\text{BinA}}^{(1)}, \beta_{\text{BinA}}^{(2)}) \quad \beta = \begin{pmatrix} 7.25 \\ -5.75 \\ 7.5 \\ -7.25 \\ 5.75 \\ -7.5 \end{pmatrix}$$

$$SSE := \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^n \left[ (Y_{\text{blocks}_{i,j}_k}) - \mu_{i,j} \right]^2$$

$$SSE = 42$$

$$n \cdot \beta^T \cdot \beta = (567.5)$$

$$e^T \cdot e = (42)$$

**β in vector form ^**

$$SSTO := \sum_{i=1}^a \sum_{j=1}^b \sum_{k=1}^n \left[ (Y_{\text{blocks}_{i,j}_k}) - \mu_{\text{dotdot}} \right]^2$$

$$SSTO = 766$$

$$(Y - \mu_{\text{dotdot}})^T \cdot (Y - \mu_{\text{dotdot}}) = (766)$$

**ANOVA Table for Nested design:**

Sum of Squares:	Degrees of Freedom:	Mean Squares:
SSA = 156.5	$df_A := a - 1$ $df_A = 2$	$MSA := \frac{SSA}{df_A}$ $MSA = 78.25$
$SSB_{inA} = 567.5$	$df_{BinA} := a \cdot (b - 1)$ $df_{BinA} = 3$	$MSB_{inA} := \frac{SSB_{inA}}{df_{BinA}}$ $MSB_{inA} = 189.1667$
SSE = 42	$df_E := a \cdot b \cdot (n - 1)$ $df_E = 6$	$MSE := \frac{SSE}{df_E}$ $MSE = 7$
SSTO = 766	$df_T := N - 1$ $df_T = 11$	

^ values verified KNNL p. 1097

**F-Test for Independent Factor Effects:****Hypotheses:**

$$H_0: \alpha_i = 0 \text{ for all } i$$

$$H_1: \text{At least one } \alpha_i \neq 0$$

**Test Statistic:**

$$F := \frac{MSA}{MSE} \quad F = 11.1786$$

**Critical Value of the Test:**

$$\alpha := 0.05 \quad < \text{Probability of Type I error must be explicitly set}$$

$$CV := qF[1 - \alpha, (a - 1), (n - 1) \cdot a \cdot b] \quad CV = 5.1433$$

$$a - 1 = 2 \\ (n - 1) \cdot a \cdot b = 6$$

**Decision Rule:**

IF  $F > C$ , THEN REJECT  $H_0$  OTHERWISE ACCEPT  $H_0$

**Probability Value:**

$$P := 1 - pF[F, a - 1, (n - 1) \cdot a \cdot b] \quad P = 0.0095$$

< test verified KNNL p. 1098

**F-Test for Nested Factor Specific Effects:****Hypotheses:**

$$H_0: \beta_{j \text{ in } i} = 0 \text{ for all values of } j \text{ within a specific level } i$$

$$H_1: \text{At least one } \beta_{j \text{ in } i} \neq 0$$

**Test Statistic:**

$$F := \frac{MSB_{inA}}{MSE} \quad F = 27.0238$$

**Critical Value of the Test:**

$$\alpha := 0.05 \quad < \text{Probability of Type I error must be explicitly set}$$

$$CV := qF[1 - \alpha, a \cdot (b - 1), (n - 1) \cdot a \cdot b] \quad CV = 4.7571$$

**Decision Rule:**

IF  $F > C$ , THEN REJECT  $H_0$  OTHERWISE ACCEPT  $H_0$

**Probability Value:**

$$P := 1 - pF[F, a \cdot (b - 1), (n - 1) \cdot a \cdot b] \quad P = 0.0007$$

< test verified KNNL p. 1098

**Prototype in R:**

```
#NESTED ANOVA
require(car)
```

```
#READ STRUCTURED DATA TABLE WITH NUMERIC CODED FACTOR
```

```
K=read.table("c:/2008LinearModelsData/TrainingSchoolR.txt")
```

```
K
```

```
attach(K)
```

```
Y=Score
```

```
A=factor(School)
```

```
B=factor(Instructor)
```

```
> K
```

	Score	School	Instructor
1	25	1	1
2	29	1	1
3	14	1	2
4	11	1	2
5	11	2	1
6	6	2	1
7	22	2	2
8	18	2	2
9	17	3	1
10	20	3	1
11	5	3	2

```
contrasts(A)=contr.sum
```

```
contrasts(B)=contr.sum
```

```
FM=lm(Y~A+B%in%A)
```

```
anova(FM)
```

```
Anova(FM, type="3")
```

```
detach(K)
```

```
> anova(FM)
```

```
Analysis of Variance Table
```

```
Response: Y
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
A	2	156.50	78.25	11.179	0.009473 **
A:B	3	567.50	189.17	27.024	0.000697 ***
Residuals	6	42.00	7.00		

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

**ANOVA tables are identical >**

```
> Anova(FM, type="3")
```

```
Anova Table (Type III tests)
```

```
Response: Y
```

	Sum Sq	Df	F value	Pr(>F)
(Intercept)	2700.0	1	385.714	1.130e-06 ***
A	156.5	2	11.179	0.009473 **
A:B	567.5	3	27.024	0.000697 ***
Residuals	42.0	6		

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

**factor A results confirmed >**

**factor B in A results confirmed >**

**KNNL p. 1098**